**Research Paper** 

# Prognostic significance of *CTNNB1* mutation in hepatocellular carcinoma: a systematic review and meta-analysis

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# ABSTRACT

Backgrounds: Hepatocellular Carcinoma (HCC) is one of the most common malignant cancers in humans and has a high fatality rate. In recent years, researchers have verified that the Wnt/ $\beta$ -catenin signaling pathway affects the clinicopathological features and prognosis of patients with HCC. Although many studies have investigated the relationship between Wnt/ $\beta$ -catenin signaling pathway and HCC, the prognostic value of  $\beta$ -catenin in HCC remains inconclusive. *CTNNB1* (Catenin Beta-1) is an important factor in the Wnt/ $\beta$ -catenin signaling pathway. However, no consensus has been reached on the clinical and prognostic significance of *CTNNB1* mutations in HCCs.

Methods: Eligible studies and relevant data were obtained from PubMed, Web of Science, Elsevier, Cochrane Library, Ovid, and Embase databases. The correlation between *CTNNB1* mutations and clinical/prognosis of patients were evaluated. A fixed- or random-effects model was used to calculate pooled odds ratios (OR) and 95% confidence intervals (CI).

Results: Seventeen studies matched the selection criteria, and 1828 patients were included. This meta-analysis demonstrated that patients with HCC with *CTNNB1* mutations had favorable clinicopathological features and survival. The combined ORs of 1-, 3- and 5-year overall survival were0.52 (n = 6 studies, 95% CI: 0.34–0.81, Z = 2.89, P = 0.004, 0.28 (n = 6 studies, 95% CI: 0.18–0.42, Z = 6.03, P < 0.00001), -0.22 (n = 6 studies, 95% CI: 0.37–0.06, Z = 2.78, P = 0.005), respectively. Additionally, *CTNNB1* mutation might be significantly associated with differentiation (OR = 0.54, 95% CI: 0.36–0.81, Z = 2.98, P = 0.003), TMN stages (Tumor, Node, Metastasis staging classification) (OR = -0.25, 95% CI:-0.33–0.18, Z = 6.60, P < 0.00001), liver cirrhosis (OR = 0.21, 95% CI:0.11–0.39, Z = 4.94, P < = 0.00001), and HBV (Hepatitis B Virus) infection (OR = 0.44, 95% CI:0.31–0.64, Z = 4.37, P < 0.0001), but not with tumor size, metastasis, vascular invasion, and HCV infection.

Conclusions: *CTNNB1* mutation can serve as an indicator of favorable prognosis as well as a novel target for treatment in HCC.

# **INTRODUCTION**

Hepatocellular carcinoma (HCC) causes nearly half a million deaths annually worldwide. Owing to the increasing incidence of hepatitis B virus (HBV) infection, HCC has become a fast-growing cancer in Asian countries, especially in China [1]. The

incidence of HCC is expected to increase in the future. However, few therapies can improve the prognosis of patients with HCC [2]. Surgery is the primary curative treatment for HCC. However, no more than 50% of patients survive longer than five years after surgery, even when diagnosed and operated at an early stage [3]. Recently, molecular

targeted therapy has brought a glimmer of hope for the treatment of HCC [4]. Owing to its favorable overall survival, the FDA approved the multi-kinase inhibitor sorafenib for the treatment of advanced HCC [5, 6]. As a targeted anticancer molecule, sorafenib has demonstrated only partial efficacy in advanced HCC. Therefore, it is mandatory to better understand the genes and signaling pathways involved in the tumorigenesis and progression of HCC and to identify more effective druggable targets for improving HCC management.

Over the past few years, Wnt pathway activation in HCC has been reported in several studies [7, 8]. This pathway has been indicated to play an important role in the clinicopathological features and prognosis of HCC [9]. According to a recent report, CTNNB1 is one of the most frequently mutated genes in HCC [10]. Mutation of CTNNB1, which is the key downstream effector of the pathway, appears to be the main cause of activation of the Wnt pathway [11]. β-catenin, a protein expressed by CTNNB1, integrates the intercellular E-cadherincatenin adhesion system, the disruption of which has been observed in HCC [12]. β-catenin is normally located in the cytomembrane and is directly connected to E-cadherin, which in turn forms an adhesion complex. This adhesion complex, which can be degraded by phosphorylation or ubiquitination, can regulate cell-cell adhesion and maintain tissue architecture and function. In HCC, unbound β-catenin translocates to the nucleus and regulates the transcription of target genes relevant to cell proliferation and cell cycle progression. β-catenin accumulation in the cytoplasm and/or nucleus is thought to be closely associated with poor prognosis and deep invasion in HCC patients, independent of tumor stage [13]. Recently, mutations of CTNNB1 have been detected in human HCC, but the clinical implications of the CTNNB1 mutation are still unclear. Our meta-analysis showed that mutant CTNNB1 was associated with favorable clinical outcomes and survival in patients with HCC.

# **MATERIALS AND METHODS**

#### Study selection

PubMed, Web of Science, Elsevier, Cochrane Library, Ovid, and Embase databases were searched for articles published until January 20, 2023. The terms used in the search were "*CTNNB1* or beta-catenin, or  $\beta$ -catenin" and "prognostic or prognosis or survival" and "hepatocellular carcinoma or HCC or liver cancer or liver tumor or hepatic cancer or liver tumor or liver neoplasms". The reference lists of all retrieved articles were manually searched. Only studies published in English were included. Two reviewers (GLL and GQS) completed the systematic literature search and extracted the following parameters from each study: study population characteristics, number of participants, sex ratio, first author, and year of publication.

#### Criteria for inclusion and exclusion

Inclusion criterion:

- (1) Patients with HCC were diagnosed by pathology;
- (2) Information about *CTNNB1* mutation, OS (Overall Survival), and other clinicopathological features were provided;
- (3) The *CTNNB1* mutation was sequenced for exon 3, SSCP analysis of exon 3, Sanger sequencing, mass array, PCR, or other methods in primary HCC tissue
- (4) The study with the highest quality assessment was enrolled when more than one study was reported by one individual author;
- (5) Studies were published in English.

Exclusion criterion:

- (1) Articles not related to the clinic;
- (2) Overlapping publications;
- (3) Information about *CTNNB1* mutation or OS or other clinicopathological features that were not clearly reported;
- (4) Abstracts, reviews, letters, editorials, and expert opinions;
- (5) Non-English publications.

#### Data extraction and literature quality assessment

Two reviewers (GLL and GQS) independently evaluated each study, and relevant characteristics were listed: (1) the first author and publication year; (2) population origin; (3) number of cases; (4) mean age, (5) gender, (6) the number of cases with *CTNNB1* mutation; (7) level of evidence, (8) disease stage, (9) clinicopathological features, (10) methods of evaluating *CTNNB1* mutation, and (11) OS data.

The quality of each study was assessed using the Newcastle-Ottawa scale (NOS), which evaluates

various aspects of the methodology, including selection, comparability, and outcome [14]. The final scores ranged from 0 (lowest) to 9 (highest); the higher the value, the better the eligibility.

#### Statistical analysis

Review Manager (RevMan) software (version 5.2; Cochrane Collaboration) was used for the metaanalysis. Odds ratios (OR) combined with 95% confidence intervals (CI) were analyzed to evaluate the association of *CTNNB1* mutation with the prognosis and clinicopathological factors of HCCs. Pooled ORs and 95%CIs were used as the recommended summary statistics. A fixed- or random-effects model was used to calculate pooled effects. Funnel plots, which were used to examine the risk of potential publication bias, were constructed using Egger's test and Begg's test. Heterogeneity was evaluated by  $I^2$ , and  $I^2$  statistics of  $\geq$ 50%, defined as heterogeneity. Statistical significance was set at P < 0.05.

#### Availability of data and materials

All data and materials were availability from the web.

#### **Consent to publication**

All co-authors consented to publish this paper.

## RESULTS

#### Selection of trials

The original search strategy retrieved 223 publications. After screening, 187 studies were excluded, and 36 papers were captured. Of these, 19 were excluded because of a lack of adequate data on *CTNNB1* mutations and specific parameters. Thus, 17 studies with sufficient evaluation met the inclusion criteria and were retrieved for further evaluation. A flowchart of the strategies is presented in Figure 1.

#### **Study characteristics**

The patient characteristics in each selected study are shown in Table 1. The total number of patients was 1828, with 319 *CTNNB1* mutations. The mean incidence of *CTNNB1* mutation was 17.5%. The information extracted from the selected studies included the gene type of *CTNNB1*, prognosis, disease stage, methods of evaluating *CTNNB1* mutation, OS data, and clinicopathological features. The ORs and 95% CI between *CTNNB1* mutation and OS are provided. All studies retrieved in this meta-analysis were performed properly, and the gene type of *CTNNB1* was determined by sequencing of exon 3, SSCP analysis of exon 3, Sanger sequencing, mass array, PCR, or other methods without subjective interference. The primary mutation was found in exon 3.



Figure 1. Flow chart of literature search strategies.

First author and year	Country /region	No. of patients	Mean age	Gender (M/F)	Mu/total	Level of evidence	Stage	Clinicopatholo gical features	Method	Provided OS data
Ding [15] 2014	China	156	53.09±11.19	NR	15/156	5	I–IV	D,T	Mass array	Yes
Lin [16] 2010	Taiwan	160	57(14-88)	122/38	22/128	5	NR	NR	Direct sequencing of exon 3	Yes
Lu [17] 2014	Taiwan	115	56.3 (23.6–83.1)	97/18	21/115	5	I–IV	T,M	Direct sequencing of exon 3	Yes
Mao [18] 2001	Taiwan	372	NR	293/162	36/372	5	I–III	D,T	PCR	Yes
Wong [19] 2001	Hong Kong	60	54(28-74)	46/14	6/60	4	I–IV	D	PCR	NR
Yuan [20] 2013	Taiwan	305	55.09 (15–88)	239/66	32/214	4	I–IV	Т	direct sequencing of exon 3	Yes
Cavard [21] 2006	France	42	NR	NR	21/42	3	NR	NR	sequencing	NR
Cieply [22] 2009	USA	25	NR	19/6	9/25	3	I–IV	T,M	Direct sequencing of exon 3	NR
Hsu [23] 2000	Japan	125	63(16-79)	88/37	57/434	3	I–IV	D,T	PCR	Yes
Huang [24] 1999	Japan/ Switzerland	16+6	NR	NR	9/22	3	NR	D	DNA sequence	NR
Kim [25] 2008	Korea	36	57.7(34-71)	32/4	1/36	3	I–IV	D,T,M	sequencing	NR
Puig [26] 2001	France	137	NR	110/27	32/137	3	NR	D,M	Direct sequencing	NR
	United States	7	56.86	3⁄4	5/7	3	I–III	Т	Sanger sequencing	NR
Li [27] 2011	China The Netherlands	1 1	68 53	1/0 1/0	1/1 0/1				1 0	
Park [28] 2005	Korea	92	51.6(26-89)	75/17	13/32	3	I–IV	D,T	SSCP analysis of exon 3	NR
Taniguchi [29] 2002	USA	73	NR	41/32	14/73	3	NR	D	PCR	NR
Tornesello [30] 2013	Italian	67	NR	53/14	10/67	3	NR	D	DNA sequence electropherograms	NR
Rossi [31] 2007	France	32	NR	NR	15/32	3	NR	NR	sequencing	NR

Table 1. Characteristics of studies included in the meta-analysis.

D, histologic differentiation degree; T, depth of tumor invasion; M, metastasis (include those with more than one Nodules); OS, overall survival; NR, not reported; M/F, male/female; Mu, *CTNNB1* mutation.

#### Quality assessment

The methodological quality of the 17 studies was assessed using NOS. On the basis of the NOS, 4 studies scored 5 points [15–18], 2 studies scored 4 [19, 20], 11 studies scored 3 [21–31] in total of the 17 studies. Studies with a score  $\geq$  5 were defined as high-quality (Table 1).

# Impact of CTNNB1 mutation on overall survival (OS)

Some of the included studies [15–18, 20, 23] provided the ORs and 95%CI directly or indirectly when discussing the correlation between *CTNNB1* mutation

and *OS*. This meta-analysis systematically assessed the association of *CTNNB1* mutation with OS in patients with HCC at 1-year, 3-year and 5-year, respectively. It was demonstrated that *CTNNB1* mutation significantly correlated with poor 1-, 3-and 5-year OS, as shown in Figure 2. And the pooled ORs were 0.52(n = 6 studies, 95% CI: 0.34–0.81, Z = 2.89, P = 0.004, 0.28 (n =6 studies, 95% CI: 0.18–0.42, Z = 6.03, P < 0.00001), -0.22(n = 6 studies, 95% CI: 0.37–0.06, Z = 2.78, P = 0.005) respectively. The statistical heterogeneity of the 1-, 3-and 5-year OS was 18%, 0%, and 81%, respectively. The above results suggest that *CTNNB1* mutation is correlated with a favorable prognosis for HCC.

# Correlation of *CTNNB1* mutation with clinicopathological parameters

This meta-analysis assessed the relationship between *CTNNB1* mutations and clinicopathological parameters, including metastasis, vascular invasion, tumor size, differentiation, TNM stages, liver cirrhosis, and HBV/HCV infection (Table 2). Ten studies [15, 18, 19, 23–26, 28–30] evaluated the correlation between *CTNNB1* mutations and differentiation (Figure 3). The pooled OR was 0.54 (95%CI: 0.36-0.81, Z=2.98, P=0.003). This result showed that there was a significant correlation between *CTNNB1* mutation and differentiation. Nine studies [15, 17, 18, 20, 22, 23, 25, 27, 28] evaluated the correlation of *CTNNB1* mutations

with TNM stages (T3/T4 versus T1/T2) (Figure 4). The pooled OR was -0.25 (95%CI: -0.33--0.18, Z=6.6, P<0.00001). This result indicated that there was a significant correlation between *CTNNB1* mutations and the TNM stages of HCC. Additionally, we also assessed the relationship between *CTNNB1* mutation and liver cirrhosis of HCC.4 studies [19, 21, 22, 26] evaluated the correlation of *CTNNB1* mutation with liver cirrhosis (Figure 5). The pooled OR was 0.21 (95%CI: 0.11-0.39, Z=4.94, P<0.00001). This result showed a significant correlation between *CTNNB1* mutations and liver cirrhosis. We found that *CTNNB1* mutation had a better effect on the above two clinicopathological features. For other parameters, such as metastasis (Supplementary Figure 1), vascular

A	CTNNB1 wi	ld-type	CTNNB1 m	utation		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
Ding 2014	98	123	12	13	7.3%	0.33 [0.04, 2.63]	
Hsu 2000	217	319	40	47	37.0%	0.37 [0.16, 0.86]	<b>_</b>
Lin 2010	91	128	20	22	16.4%	0.25 [0.05, 1.11]	
Lu 2014	26	79	5	21	8.8%	1.57 [0.52, 4.76]	
Mao 2001	29	35	26	28	8.2%	0.37 [0.07, 2.01]	
Yuan 2013	122	182	24	32	22.3%	0.68 [0.29, 1.60]	-•+
Total (95% CI)		866		163	100.0%	0.52 [0.34, 0.81]	•
Total events	583		127				
Heterogeneity: Chi <sup>2</sup> =	: 6.09, df = 5 (F	<sup>o</sup> = 0.30);	I <b>²</b> = 18%				
Test for overall effect	: Z = 2.89 (P =	0.004)					0.01 0.1 1 10 10 Favours [CTNNB1 Mu] Favours [CTNNB1 Wt]

	CTNNB1 wi	ld-type	CTNNB1 m	utation		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% CI
Ding 2014	96	123	12	13	5.1%	0.30 [0.04, 2.38]	
Hsu 2000	140	319	34	47	35.7%	0.30 [0.15, 0.59]	
Lin 2010	54	128	17	22	18.0%	0.21 [0.07, 0.62]	<b>_</b>
Lu 2014	4	79	1	21	1.6%	1.07 [0.11, 10.08]	
Mao 2001	15	35	23	28	15.7%	0.16 [0.05, 0.53]	
Yuan 2013	74	182	22	32	23.9%	0.31 [0.14, 0.70]	
Total (95% CI)		866		163	100.0%	0.28 [0.18, 0.42]	◆
Total events	383		109				
Heterogeneity: Chi <sup>2</sup> :	= 2.52, df = 5 (F	<sup>o</sup> = 0.77);	I <sup>z</sup> = 0%				
Test for overall effec	t: Z = 6.11 (P <	0.00001)	)				0.01 0.1 1 10 10 Favours [CTNNB1 Mu] Favours [CTNNB1 Wt]

	CTNNB1 wil	d-type	CTNNB1 mu	tation		Risk Difference	Risk Difference
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
Ding 2014	95	123	12	13	17.2%	-0.15 [-0.31, 0.01]	
Hsu 2000	89	319	29	47	17.8%	-0.34 [-0.49, -0.19]	_ <b></b>
Lin 2010	40	128	12	22	14.7%	-0.23 [-0.46, -0.01]	
Lu 2014	3	79	1	21	19.5%	-0.01 [-0.11, 0.09]	-+-
Mao 2001	13	35	21	28	14.5%	-0.38 [-0.61, -0.15]	
Yuan 2013	51	182	17	32	16.3%	-0.25 [-0.44, -0.07]	
Total (95% CI)		866		163	100.0%	-0.22 [-0.37, -0.06]	•
Total events	291		92				
Heterogeneity: Tau <sup>2</sup> :	= 0.03; Chi <sup>z</sup> = 2	6.29, df:	= 5 (P < 0.000	1); I <sup>2</sup> = 8	1%		
Test for overall effect	: Z= 2.78 (P=	0.005)					-1 -0.5 0 0.5 Favours (CTNNB1 Mu) Favours (CTNNB1 Wt)

Figure 2. Forest plot of odds ratio for the association of CTNNB1 mutation with 1-year (A), 3-year (B) and 5-year (C) overall survival.

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Outcome of interest	No. of studies	Number of tissue samples	OR/WMD	95% CI	P value	I <sup>2</sup> (%)
Overall Survival [15–18, 20, 23]						
1 year	6	<i>CTNNB1</i> Mu=163, <i>CTNNB1</i> Wt =866	0.52	0.34-0.81	0.004	18
3 year	6	<i>CTNNB1</i> Mu=163, <i>CTNNB1</i> Wt =866	0.28	0.18-0.42	< 0.00001	0
5 year	6	<i>CTNNB1</i> Mu=163, <i>CTNNB1</i> Wt =866	-0.22	-0.370.06	0.005	81
Differentiation grade [15, 18, 19 23–26, 28–30,]	10	<i>CTNNB1</i> Mu=184, <i>CTNNB1</i> Wt =680	0.54	0.36-0.81	0.003	0
TMN stage [15, 17, 18, 20, 22, 23, 25, 27, 28]	9	<i>CTNNB1</i> Mu=174, <i>CTNNB1</i> Wt =841	-0.25	-0.330.18	< 0.00001	39
Metastasis [17, 22, 25, 26]	4	<i>CTNNB1</i> Mu=69, <i>CTNNB1</i> Wt =242	1.25	0.93-1.66	0.14	0
Vascular invasion [17, 22, 26]	3	<i>CTNNB1</i> Mu=70, <i>CTNNB1</i> Wt =223	1.42	0.82-2.45	0.21	43
Liver cirrhosis [19, 21, 22, 26]	4	<i>CTNNB1</i> Mu=76, <i>CTNNB1</i> Wt =138	0.21	0.11-0.39	< 0.00001	0
Tumor size [19, 22, 25, 26, 28]	5	<i>CTNNB1</i> Mu=67, <i>CTNNB1</i> Wt =171	1.24	0.37-4.11	0.72	54
HBV [15, 17–19, 21, 23, 25–28, 30, 31]	12	<i>CTNNB1</i> Mu=213, <i>CTNNB1</i> Wt =795	0.44	0.31-0.64	< 0.0001	0
HCV [15, 17, 21, 22, 25, 27, 28, 30, 31]	9	<i>CTNNB1</i> Mu=111, <i>CTNNB1</i> Wt =356	1.70	0.93-3.11	0.09	0

Table 2. Meta-analysis comparing HCC with *CTNNB1* mutation and wild-type.

OR, odds ratio; WMD, weighted mean difference; CI, confidence interval; Mu, mutation; Wt, wild-type.

invasion (Supplementary Figure 2), and tumor size (Supplementary Figure 3), of HCC showing *CTNNB1* mutation, the pooled ORs were1.25(n=4studies, 95%CI:0.93-1.66, Z=1.49, *P*=0.14),1.42 (n=3 studies,

95% CI:0.82-2.45, Z=1.26, P=0.21) and 1.24 (n=5 studies, 95% CI:0.37-4.11, Z=0.35, P=0.72), respectively, demonstrating that *CTNNB1* mutation had no significant correlation with these parameters.

	CTNNB1 wil	d-type	CTNNB1 m	utation		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
Ding 2014	18	141	5	15	11.7%	0.29 [0.09, 0.95]	
Hsu 2000	200	319	35	47	33.7%	0.58 [0.29, 1.15]	
Huang 1999	11	13	9	9	2.9%	0.24 [0.01, 5.68]	
Kim 2008	25	34	1	1	1.1%	0.89 [0.03, 23.93]	
Mao 2001	26	36	24	30	10.8%	0.65 [0.20, 2.06]	
Park 2005	15	19	12	13	4.4%	0.31 [0.03, 3.18]	
Puig 2001	43	98	25	38	30.0%	0.41 [0.19, 0.89]	
Taniguchi 2002	4	6	12	14	3.6%	0.33 [0.03, 3.20]	
Tornesello 2013	10	11	6	10	0.8%	6.67 [0.60, 74.51]	
Wong 2001	2	3	3	7	0.9%	2.67 [0.16, 45.14]	
Total (95% CI)		680		184	100.0%	0.54 [0.36, 0.81]	•
Total events	354		132				
Heterogeneity: Chi <sup>2</sup> =	7.80, df = 9 (P	= 0.55);	I² = 0%				
Test for overall effect							0.01 0.1 1 10 100 Favours [CTNNB1 Mu] Favours [CTNNB1 Wt]

#### Figure 3. Forest plot of odds ratio for the association of *CTNNB1* mutation with differentiation grade.

#### Correlation of CTNNB1 mutation with etiology

In this meta-analysis, the correlation between *CTNNB1* mutations and etiology (HBV/HCV infection) was

evaluated. As shown in Figure 6, 12 studies [15, 17, 18, 19, 21, 23, 25–28, 30, 31] assessed the relationship between *CTNNB1* mutations and HBV. The combined ORs were 0.44(95%CI:0.31-0.64, Z=4.37, P<0.0001),

	CTNNB1 wil	d-type	CTNNB1 m	utation		Risk Difference	Risk Difference
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
Cieply 2009	10	14	6	9	4.2%	0.05 [-0.34, 0.44]	
Ding 2014	93	141	14	15	10.4%	-0.27 [-0.42, -0.13]	_ <b>-</b> _
Hsu 2000	129	319	34	47	31.3%	-0.32 [-0.46, -0.18]	
Kim 2008	25	34	1	1	0.7%	-0.26 [-0.88, 0.35]	
Li 2011	2	2	5	6	1.1%	0.17 [-0.35, 0.69]	
Lu 2014	28	94	7	21	13.1%	-0.04 [-0.26, 0.19]	
Mao 2001	23	36	28	30	12.5%	-0.29 [-0.47, -0.11]	_ <b></b>
Park 2005	12	19	9	13	5.9%	-0.06 [-0.39, 0.27]	
Yuan 2013	55	182	22	32	20.8%	-0.39 [-0.56, -0.21]	
Total (95% CI)		841		174	100.0%	-0.25 [-0.33, -0.18]	◆
Total events	377		126				
Heterogeneity: Chi <sup>2</sup> =	= 13.17, df = 8 (	(P = 0.11)	); I² = 39%				
Test for overall effect	: Z = 6.60 (P <	0.00001)	•				-1 -0.5 0 0.5 1 Favours [ <i>CTNNB1</i> Mu] Favours [ <i>CTNNB1</i> Wt]



	CTNNB1 wild	i-type	CTNNB1 mu	ation		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI	M-H, Fixed, 95% CI
Cavard 2008	7	21	11	21	17.2%	0.45 (0.13, 1.58)	
Cieply 2009	5	16	6	9	12.4%	0.23 [0.04, 1.30]	
Puig 2001	18	98	24	39	65.7%	0.14 [0.06, 0.32]	
Wong 2001	0	3	3	7	4.8%	0.18 (0.01, 4.86)	· · · · · · · · · · · · · · · · · · ·
Total (95% CI)		138		76	100.0%	0.21 [0.11, 0.39]	◆
Total events	30		44				
Heterogeneity: Chi*=	2.39, df = 3 (P	= 0.50);	l* = 0%				
Test for overall effect	Z=4.94 (P < 0	.00001)					0.01 0.1 1 10 100 Favours [CTNNB1 Mu] Favours [CTNNB1 Wi]

#### Figure 5. Forest plot of odds ratio for the association of CTNNB1 mutation with liver cirrhosis.

	CTNNB1 wil	d-type	CTNNB1 m	utation		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
Cavard 2006	12	21	18	21	9.0%	0.22 [0.05, 0.99]	
Ding 2014	43	141	3	15	4.4%	1.76 [0.47, 6.54]	
Hsu 2000	91	319	25	47	36.4%	0.35 [0.19, 0.65]	
Kim 2008	14	34	1	1	1.9%	0.24 [0.01, 6.21]	←
Li 2011	3	3	5	6	0.6%	1.91 [0.06, 61.34]	
Lu 2014	29	94	8	21	10.6%	0.72 [0.27, 1.94]	
Mao 2001	11	36	16	30	14.2%	0.39 [0.14, 1.06]	
Park 2005	1	19	0	13	0.6%	2.19 [0.08, 57.98]	
Puig 2001	71	97	24	27	11.8%	0.34 [0.09, 1.23]	
Rossi 2007	11	17	14	15	6.1%	0.13 [0.01, 1.25]	
Tornesello 2013	9	11	10	10	2.7%	0.18 [0.01, 4.27]	•
Wong 2001	0	3	2	7	1.7%	0.31 [0.01, 8.68]	
Total (95% Cl)		795		213	100.0%	0.44 [0.31, 0.64]	•
Total events	295		126				
Heterogeneity: Chi <sup>2</sup> =	9.96, df = 11 (	P = 0.53	); I² = 0%				
Test for overall effect:	:Z=4.37 (P < )	0.0001)					0.01 0.1 1 10 100 Favours [ <i>CTNNB1</i> Mu] Favours [ <i>CTNNB1</i> Wt]

#### Figure 6. Forest plot of odds ratio for the association of CTNNB1 mutation with etiology (HBV).

with no significant statistical heterogeneity ( $I^2=0\%$ ). This result indicates that a better effect was observed between *CTNNB1* mutation and HBV infection. However, as shown in Supplementary Figure 4, 9 studies [15, 17, 21, 22, 25, 27, 28, 30, 31] assessed the relationship between *CTNNB1* mutations and HCV. The combined ORs were 1.70(95%CI: 0.93-3.11, Z=1.72, P=0.09).

### **Publication bias**

For the studies included in this meta-analysis, Begg's test indicated that there was no significant publication bias after assessing the funnel plot (Supplementary Figures 5-15).

# **DISCUSSION**

Human HCCs with activation of the Wnt/β-catenin pathway demonstrate unique gene expression patterns and pathological features. Activated Wnt/β-catenin synergizes with multiple signaling cascades to drive HCC formation, and it functions through its downstream effectors [32]. The aberrant Wnt pathway in HCC, has been well-studied and proven to be involved in the prognosis of HCC [33]. β-catenin is a key downstream effector of the Wnt signaling pathway and plays a crucial role. β-catenin has been the focus of attention as an attractive therapeutic target [9]. According to our previous observations, cytoplasmic and/or nuclear expression of Bcatenin could serve as a potential predictor for the progression and prognosis of patients with HCC and act as a novel target for the developed therapies [13]. *CTNNB1*, the coding gene of  $\beta$ -catenin, is one of the most frequently mutated genes in HCC [10, 34]. Hsu et al. first reported a relationship between β-catenin mutations and prognosis in patients with HCC [23]. Based on the resected HCC of patients, their report showed that CTNNB1 mutations were associated with tumor differentiation. HBV infection, and clinical prognosis. However, according to a study by Lu et al., there is no prognostic significance of CTNNB1 mutation in patients with HCC [17]. Moreover, no difference in CTNNB1 mutation rate was observed between patients with HCC with or without HBV infection. Tumor mutational burden (TMB) was verified to be closely associated with immune checkpoint inhibitors, but it is unclear whether gene mutation has an effect on immunotherapy of HCC. Mo et al. firstly revealed the underlying association between CTNNB1 mutation and immunotherapy, and they speculated that CTNNB1 mutation may modulate NK cells by affecting CD96 [35]. To date, however, tremendous work has been done to investigate the association between CTNNB1 mutations and clinicopathological characteristics and prognosis of patients with HCC; however, no conclusive results have been achieved. Wang et al. reported that the *CTNNB1* mutation was correlated with a favorable prognosis in HCC in a meta-analysis [36]. Of note, the studies enrolled in their meta-analysis failed to provide adequate information about the relationship between *CTNNB1* mutations and clinical prognosis. The current study focused on *CTNNB1* mutations in HCC in a clinical setting. We showed that *CTNNB1* mutation has a robust effect on the clinical and prognosis of patients with HCC.

A number of studies have implicated an alteration in CTNNB1, including gene mutation and protein overexpression, in HCC [34, 37]. CTNNB1 plays a crucial role in hepatocyte adhesion and Wnt signaling pathway [38]. CTNNB1 is mainly located in the cell membrane. By binding to the lymphoid enhancer factor (LEF)/T-cell factor (TCF) family of DNA binding proteins, CTNNB1 enters the nucleus and regulates transcription of target genes such as c-mycor cyclin D1, resulting in proliferation and metastasis of liver tumor cells [39]. Xiao X et al. demonstrated that CTNNB1 mutation in HCC led to a decrease in chemokine expression and subsequent suppression of immune cell infiltration [40]. Recently, somatic mutations in CTNNB1 have been demonstrated not only in animal models but also in human HCC [10]. Comprehensive analysis of clinical samples has identified immunological and molecular classification of HCC, and the CTNNB1-mutated subtype exhibits characteristics of immunosuppressive distinctive tumor microenvironment [41]. However, the clinical implications of CTNNB1 mutation in human HCC are unclear [12, 42]. Paradoxical data exist concerning the prognostic value and clinicopathological significance of cytoplasmic and/or nuclear CTNNB1 accumulation. These discrepancies are most likely due to the gene type of CTNNB1. Cytoplasm and/or nucleusCTNNB1 accumulation can be associated with mutations in the CTNNB1 gene and other components of the signaling pathway [43]. The most common mechanism of CTNNB1 accumulation in HCC is mutations in CTNNB1 [44]. Mutations or wild-type CTNNB1 may influence the subcellular localization of β-catenin. Most studies on CTNNB1 mutations have focused on the consensus sequence for GSK-3<sup>β</sup> phosphorylation in exon 3 and the inactivation of APC and other factors [45]. Cui et al. found that mutation of exon3 of CTNNB1 is one of the most important factors activating the abnormal Wnt signaling pathway in HCC [46]. CTNNB1 mutations and nuclear overexpression may play a key role in HCC in Chinese people. And targeting the Wnt-beta-catenin pathway may represent a valid treatment option for Chinese HCC patients [15, 47, 48].

In the past 5 years, immune-checkpoint inhibitors have revolutionized the management of HCC [49]. However, several studies demonstrated that mutation of CTNNB1 have also been associated with scarcity of immune cells in the tumor microenvironment and poor clinical response to immune checkpoint inhibitor therapy [50]. Ogawa K et al. verified that Gain-of-function Inclusion mutation of CTNNB1 contributes to resistance of ICI monotherapy through the framework of non-T-cellinflamed tumor microenvironment. Of note, the treatment effect of Atezolizumab plus bevacizumab in patients with HCC with MT CTNNB1 was comparable to those patients with WT CTNNB1. These results further implicate that bevacizumab added to Atezolizumab might improve immunosuppressive tumor microenvironment caused by CTNNB1 mutation [51]. In addition, Chen et al. proved that CTNNB1 alternation is a potential biomarker for immunotherapy prognosis in patients with HCC [52]. However, this meta-analysis concluded that HCC patients with CTNNB1 mutations appeared to have a favorable survival in comparison with wild-type CTNNB1 HCC. Additionally, CTNNB1 mutations were significantly associated with the differentiation grade, TNM stage, liver cirrhosis, and HBV infection. The reason may be closely related to the concept of CTNNB1 mutations in this meta-analysis refer to loss-of-function mutation.

However, our study has some limitations. Firstly, the clinical data used in this study were acquired from a relatively small cohort in each study, so selection bias or potential biases related to imbalanced clinical characteristics is inevitable. This will also lead to heterogeneity, which is a potential problem that may affect the results of all meta-analyses. In this study, significant heterogeneity was found when discussing the relationship between CTNNB1 mutations and 5-year overall survival and tumor size in the selected studies. Unfortunately, due to limited information, a metaregression analysis could not be conducted. To eliminate variations across the included studies, we used a random effects model. Although this method may not completely eliminate the effects of heterogeneity, its adverse effects must be weakened. The second was publication bias, which can be seen in the publication bias evaluation. As is well known, this bias was unavoidable because positive results were more likely to be published than negative ones. Moreover, most studies included in this metaanalysis failed to elucidate the relationship between CTNNB1 expression and CTNNB1 mutation. Some studies provided the number of tissue samples with CTNNB1 expression, but failed to provide information on the gene type of CTNNB1, and vice versa. What's more, the method of detecting CTNNB1 mutations was not standardized from author to author. Direct sequencing of exon 3, mass array, DNA sequence electropherograms, Sanger sequencing, SSCP analysis of exon 3, and PCR were performed. Therefore, a more standardized analysis should be performed and more prospective works and experimental clinical research should be conducted. Last but not the least, how *CTNNB1* mutation effect clinical outcomes and survival of patients with HCC and how it serves as a valuable prognostic predictor is indeterminacy. Further research on the regulation of *CTNNB1* expression by the *CTNNB1* mutation-related signaling pathway is needed and will help to elucidate the new mechanism of drug resistance, providing a theoretical basis for the prediction of drug sensitivity in HCC and the development and application of new therapeutic targets for reversing drug resistance.

In conclusion, *CTNNB1* mutation could serve as a potential predictor for the clinical and prognosis of HCC patients and act as a novel useful biomarker of molecular targeted therapies for HCC.

## Abbreviations

HCC: Hepatocellular Carcinoma; OR: Odds ratio; HBV: hepatitis B virus; OS: overall survival; *CTNNB1*: Catenin Beta-1; TMN stages: Tumor, Node, Metastasis staging classification; NOS: the Newcastle-Ottawa scale; CI: confidence intervals; PCR: Polymerase Chain Reaction.

# **AUTHOR CONTRIBUTIONS**

GLL designed this study. GLL and GQS completed the study selection and data extraction, JL completed the work of literature quality assessment, MC and GLL mainly focused on statistical analysis, and GLL and GQS wrote the paper.

# **CONFLICTS OF INTEREST**

The authors declare that they have no conflicts of interest.

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# SUPPLEMENTARY MATERIALS

# **Supplementary Figures**

	CTNNB1 wil	d-type	CTNNB1 mu	tation		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI	M-H, Fixed, 95% CI
Cieply 2009	10	16	6	9	17.1%	0.94 (0.52, 1.70)	-
Kim 2008	29	34	1	1	6.3%	1.12 [0.50, 2.53]	-
Lu 2014	28	94	6	21	21.9%	1.04 [0.50, 2.19]	
Puig 2001	63	98	17	38	54.7%	1.44 (0.98, 2.11)	-
Total (95% CI)		242		69	100.0%	1.25 [0.93, 1.66]	•
Total events	130		30				
Heterogeneity: Chi*=	1.68, df = 3 (P	= 0.64);	l² = 0%				
Test for overall effect	Z=1.49 (P=0	0.14)					0.01 0.1 1 10 100 Favours [CTNNB1 Mu] Favours [CTNNB1 Wt]

#### Supplementary Figure 1. Forest plot of odds ratio for the association of *CTNNB1* mutation with metastasis.

	CTNNB1 wil	d-type	CTNNB1 m	utation		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% CI
Cieply 2009	10	16	2	9	4.4%	5.83 (0.90, 37.82)	
Lu 2014	55	115	11	21	45.0%	0.83 [0.33, 2.11]	
Puig 2001	50	92	20	40	50.8%	1.56 [0.74, 3.29]	+
Total (95% CI)		223		70	100.0%	1.42 [0.82, 2.45]	◆
Total events	121		33				
Heterogeneity: Chi* =	3.51, df = 2 (P	= 0.17);	l≊= 43%				
Test for overall effect	Z=1.26 (P=	0.21)					0.01 0.1 1 10 100 Favours [CTNNB1 Mu] Favours [CTNNB1 WI]

#### Supplementary Figure 2. Forest plot of odds ratio for the association of *CTNNB1* mutation with vascular invasion.

	CTNNB1 will	d-type	CTNNB1 mu	tation		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
Cieply 2009	13	16	5	9	21.2%	3.47 [0.56, 21.35]	
Kim 2008	28	34	1	1	9.9%	1.46 [0.05, 40.11]	
Park 2005	9	19	10	13	24.1%	0.27 [0.06, 1.30]	
Puig 2001	56	99	12	38	35.0%	2.82 [1.28, 6.22]	<b> </b> − <b>∎</b> −
Wong 2001	0	3	2	6	9.8%	0.26 [0.01, 7.27]	• • •
Total (95% CI)		171		67	100.0%	1.24 [0.37, 4.11]	
Total events	106		30				
Heterogeneity: Tau <sup>a</sup> =	: 0.90; Chi <sup>a</sup> = 8	.68, df=	4 (P = 0.07); P	°= 54%			
Test for overall effect							0.01 0.1 1 10 100 Favours [CTNNB1 Mu] Favours [CTNNB1 WI]

Supplementary Figure 3. Forest plot of odds ratio for the association of *CTNNB1* mutation with tumor size.

	CTNNB1 wild-type		CTNNB1 mutation		Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl
Cavard 2006	16	21	12	21	18.4%	2.40 [0.64, 9.03]	
Cieply 2009	13	16	6	9	9.3%	2.17 [0.33, 14.06]	
Ding 2014	140	141	15	15	1.9%	3.02 [0.12, 77.41]	
Kim 2008	30	34	1	1	2.3%	2.26 [0.08, 64.40]	
Li 2011	1	3	2	6	5.7%	1.00 [0.05, 18.91]	
Lu 2014	75	94	16	21	34.0%	1.23 [0.40, 3.79]	
Park 2005	18	19	13	13	7.7%	0.46 [0.02, 12.10]	
Rossi 2007	12	17	9	15	18.1%	1.60 [0.37, 6.95]	
Tornesello 2013	2	11	0	10	2.7%	5.53 [0.23, 130.34]	
Total (95% CI)		356		111	100.0%	1.70 [0.93, 3.11]	◆
Total events	307		74				
Heterogeneity: Chi <sup>2</sup> = 2.07, df = 8 (P = 0.98); l <sup>2</sup> = 0%							
Test for overall effect: Z = 1.72 (P = 0.09)     0.01     0.1     1     10     100       Test for overall effect: Z = 1.72 (P = 0.09)     Favours [CTNNB1 Mu]     Favours [CTNNB1 Mu]     Favours [CTNNB1 Mu]							

Supplementary Figure 4. Forest plot of odds ratio for the association of CTNNB1 mutation with etiology (HCV).



**Supplementary Figure 5. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and 1-year overall survival (OS) in the meta-analysis.



**Supplementary Figure 6. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and 3-year overall survival (OS) in the meta-analysis.



**Supplementary Figure 7. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and 5-year overall survival (OS) in the meta-analysis.



**Supplementary Figure 8. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and differentiation grade in the meta-analysis.



**Supplementary Figure 9. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and TNM stages in the meta-analysis.



**Supplementary Figure 10. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and liver cirrhosis in the meta-analysis.



**Supplementary Figure 11. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and etiology (HBV) in the meta-analysis.



**Supplementary Figure 12. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and metastasis in the meta-analysis.



**Supplementary Figure 13. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and vascular invasion in the meta-analysis.



**Supplementary Figure 14. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and tumor size in the meta-analysis.



**Supplementary Figure 15. Funnel plot to assess publication bias.** Begg's publication bias plot showed no publication bias for studies regarding *CTNNB1* mutation and HCV in the meta-analysis.